Is Discriminability a Requirement for Reactivity? Comparing the Effects of Mixed vs. Pure List Presentations on Judgment of Learning Reactivity

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**Author Note**

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Abstract

Having participants provide judgments of learning (JOLs) at study tends to produce reactive effects on recall of cue-target word pairs. This reactivity generally produces memory improvements (i.e., positive reactivity) but only for related word pairs. For unrelated pairs, reactivity is typically not observed. Researchers have primarily investigated reactivity using study lists which contain at least two distinct pair types (i.e., related vs. unrelated pairs). Using these mixed lists, reactivity may occur because participants use pair information to inform their study goals (i.e., prioritizing related vs. unrelated pairs). The present study examined whether detection of separate pair types within mixed lists is a requisite for reactivity to occur. Experiment 1 replicated previous work showing that in mixed lists, JOLs produced positive reactivity on related pairs but are not reactive on unrelated pairs. Importantly, Experiment 1 also found that these patterns extended to pure lists, in which only one pair type is presented. Experiments 2 and 3 then extended these patterns to backward and symmetrical paired associates. Finally, across experiments, reactivity patterns reported for JOLs extended to frequency of co-occurrence judgments across pair and list types. Our findings that reactivity patterns consistently extend to pure lists supports a cue-strengthening account of reactivity.

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Is Discriminability a Requirement for Reactivity? Comparing the Effects of Mixed vs. Pure List Presentations on Judgment of Learning Reactivity

Judgments of learning (JOLs) are used to assess metamemory processes participants engage in at encoding. While JOLs can be elicited for various types of study material (e.g., text passages, Townsend & Heit, 2011; sentences, Luna, Albuquerque, & Martín-Luengo, 2019; etc.), participants commonly study cue-target pairs (e.g., word pairs like cat-dog) and are instructed to estimate their likelihood of correctly recalling the target (e.g., dog) at test if shown the cue (e.g., cat). While JOLs are used to gauge metacognitive processes (see Rhodes, 2016 for review), a growing body of research suggests that these judgments are *reactive* on learning, particularly for cue-target pair learning (e.g., Janes, Rivers, & Dunlosky, 2018; Maxwell & Huff, in press; Soderstrom, Clark, Halamish, & Bjork, 2015). Reactivity occurs whenever a measure causes participants to attend to information they might otherwise ignore, leading to changes in performance (Ericsson & Simon, 1993). Regarding JOLs, reactivity produces memory changes, which manifest as either memory benefits (i.e., *positive reactivety*) or costs (i.e., *negative reactivity*). Testing for these memory changes is simple and merely requires comparing recall for participants who make JOLs at encoding to a no-JOL control group (e.g., silent reading). However, this comparison group has commonly been omitted, particularly for studies in which JOLs are made immediately following study, as researchers have often been more interested in factors affecting the accuracy of JOLs (e.g., associative direction; Koriat & Bjork, 2005; Maxwell & Huff, 2021; multiple study trials; Koriat, Sheffer, & Ma’ayan, 2002; Meeter & Nelson, 2003) rather than the direct effects of these judgments on memory.

Although JOL studies have commonly omitted control group comparisons, interest in the potential effects of these judgments on memory is not new. Research suggests that JOLs made following a delay can produce a memory benefit (e.g., Akdoğan, Izaute, Danion, Vidailhet, & Bacon, 2016; Spellman & Bjork, 1992; see Rhodes & Tauber, 2011). However, researchers have only recently begun to explore whether concurrent and immediate JOLs (i.e., those elicited at or immediately following encoding) are similarly reactive. This is surprising, given reactive effects of immediate JOLs were reported over 50 years ago by Arbuckle and Cuddy (1969). In their early work, Arbuckle and Cuddy compared recall between two groups of participants: Those who made JOLs at study and also provided confidence judgments at test and those who only silently read each item at study but still made confidence judgments at test. This design allowed for a comparison of recall rates between participants making JOLs at encoding to a group of participants who engaged in silent reading. Overall, the authors found that JOLs produced positive reactivity. However, because all participants also made confidence judgments at retrieval, it was unclear whether providing these judgments at test was also a requirement for reactivity to occur.

More recently, Soderstrom et al. (2015) tested for reactivity by comparing recall between participants who made JOLs immediately following encoding to a silent reading control group. Across groups, participants studied cue-target word pairs, in which half of which were related (e.g., mouse-cheese) and the other half were unrelated (e.g., dog-bread). Following the JOL/study phase, participants completed a cued-recall test, which, importantly, did not require participants to make additional metacognitive judgments (cf. Arbuckle & Cuddy, 1969). Overall, Soderstrom et al. demonstrated a positive reactivity pattern in which cued-recall performance was greater for participants who made JOLs relative to the control group. However, this pattern was moderated by pair relatedness, as only related pairs showed positive reactivity. When pairs were unrelated, no differences in recall were detected between the two groups. Subsequent studies by Janes et al. (2018) and Maxwell and Huff (in press) replicated this pattern using immediate and concurrent JOLs, respectively, with both studies similarly showing that JOLs produce positive reactivity selectively on related cued-target pairs.

Although recent studies show that immediate JOLs produce positive reactivity on related pairs but have no effect on unrelated related pairs (e.g., Janes et al., 2018; Maxwell & Huff, in press; Soderstrom et al., 2015), Mitchum, Kelley, & Fox (2016) reported a different pattern. Specifically, they found no reactivity for related pairs, and JOLs for unrelated pairs showed a negative reactivity pattern relative to no-JOL control pairs. To date, it is unclear why this pattern changed as similar methodologies were used relative to other studies (e.g., Maxwell & Huff, in press; Soderstrom et al., 2015), however a meta-analysis by Double, Birney, Walker (2018) reported positive reactivity for related pairs and no reactivity for unrelated pairs across the 17 experiments they analyzed.

**Theories of JOL Reactivity**

While several theories to explain JOL reactivity have been proposed, the two most prominent accounts are the *changed-goal hypothesis* (Mitchum et al., 2016) and the *cue-strengthening account* (Soderstrom et al., 2015). First, Mitchum et al.’s changed-goal hypothesis proposes that reactivity occurs because participants shift study goals as they progress through a study list. According to this account, participants initially approach a study task with a broad goal of mastering the entire list. However, when instructed to make JOLs at study, participants realize that not all pairs will be remembered equally well, particularly when study lists contain a mix of easy and difficult pair types (i.e., related vs. unrelated pairs). As a result, participants use perceptions of item difficulty to alter their study strategies, prioritizing pairs perceived as easy at the expense of more difficult pairs. Thus, the changed-goals hypothesis predicts positive reactivity for pairs perceived as easy to learn (e.g., related pairs) and negative reactivity for pairs perceived as difficult (e.g., unrelated pairs). However, because this account is dependent on a comparison process, the changed-goal hypothesis assumes that study lists will contain at least two discernable pair types (i.e., related vs. unrelated pairs). Reactivity would not be expected to occur when lists contain only one pair type (e.g., only related or unrelated pairs).

Alternatively, Soderstrom et al.’s cue-strengthening account proposes that the act of making JOLs directs participants’ attention towards intrinsic cues about each study pair that participants use to inform their JOLs (e.g., pair relatedness; see Koriat, 1997). According to this account, reactivity occurs anytime these cues emphasized by JOLs at encoding are made available at test (e.g., cued-recall testing). As a result, positive reactivity should occur on related pairs, but no reactivity for unrelated pairs, given this pair type’s lack of relatedness cues. Furthermore, the cue strengthening account makes no predictions regarding list composition, as reactivity in this account does not require an easy/difficult comparison, just that there is the availability of intrinsic cues that could direct attentional processes at encoding.

Prior research has generally supported a cue-strengthening account over a changed-goal account. For example, Rivers, Janes, and Dunlosky (2021) replicated previous reactivity patterns reported by Janes et al. (2018) and Soderstrom et al. (2015) while also extending them to a within-subject manipulation in which participants made JOLs for only some items in a study list. Importantly, Rivers et al. showed no evidence of negative reactivity on unrelated pairs. Additionally, Myers, Rhodes, and Hausman (2020) demonstrated that positive reactivity on related pairs was contingent upon relatedness cues being made available at test, as positive reactivity on related pairs occurred when participants were tested via cued-recall and recognition but not free-recall in which relatedness cues are absent. Finally, Maxwell and Huff (in press) showed that positive reactivity on related pairs was not limited to JOLs and extended to other, non-metacognitive judgment tasks that similarly emphasize relatedness cues, including judgments of associative memory (JAMs; Maki, 2007; Valentine & Buchanan, 2013) and frequency of co-occurrence judgments. Thus, reactivity occurs whenever the judgment task emphasizes cues used at retrieval, rather than an adjustment of study goals.

**Mixed- vs. Pure-List Designs**

With few exceptions, studies investigating JOL reactivity have done so using mixed-list designs in which participants are presented with both related and unrelated pairs at study. A mixed-list design is central to the changed-goal hypothesis, as this account states that participants’ ability to discriminate between different pair types drives a shift in goals and, by extension, reactivity. Thus, this hypothesis predicts that reactivity could only occur when a mixed-list design is used, as this “easy-difficult” comparison cannot occur in a pure list in which there is only one pair type. Regarding the cue-strengthening account, however, reactivity would be expected to occur whenever the encoding task emphasizes cues used at retrieval, regardless of whether pairs are presented using mixed or pure lists. Therefore, our inclusion of pure lists in the present study provided a method to directly test these competing accounts.

Although studies investigating reactivity effects have generally used mixed-list designs, both Janes et al. (2018) and Tauber and Witherby (2019) each included pure-group comparisons. First, Janes et al.’s (2018) Experiment 2 compared JOL reactivity effects for mixed- vs. pure-list designs by having participants study 1) mixed lists of forward associative pairs and unrelated pairs, 2) pure lists of forward associative pairs, or 3) pure lists of unrelated pairs. Overall, the authors found that positive reactivity patterns typically found for mixed lists with related pairs failed to emerge when a pure list was used, suggesting that reactivity effects were contingent on participants being able to discriminate between different pair types. In contrast, Tauber and Witherby (2019) showed a reactivity effect for forward pairs presented using a pure list. However, because Tauber and Witherby only used pure related lists and did not include a mixed list comparison group, it remains unclear how these observed reactivity effects compare to a mixed list (i.e., whether reactivity effects would be greater when using a mixed list relative to a pure list) or whether this effect would also extend to a pure list of unrelated pairs.

Given these discrepancies, the goal of the present study was to provide a more direct test of list type on reactivity by comparing recall for a group of participants who studied mixed lists which contained related and unrelated pairs to separate groups of participants who studied either pure lists of only related or unrelated word pairs. First, Experiment 1 provided a direct replication of Janes et al.’s (2018) second experiment by comparing reactivity effects for forward and unrelated pairs across mixed and pure lists. Experiments 2 and 3 then expanded upon Experiment 1 by comparing unrelated pairs to backward and symmetrical pairs, respectively. Each experiment, therefore, provided three separate tests of list effects on reactivity. Finally, because Maxwell and Huff (in press) showed that reactivity effects extend to other, non-metacognitive judgment tasks, each experiment included an additional frequency judgment comparison group in which participants rated the likelihood that paired items would appear together in everyday language rather than making a JOL. This additional comparison was included to 1) test whether the reactivity effects for frequency judgments initially reported by Maxwell and Huff would replicate for mixed groups and 2) test whether these judgments would continue mirror JOL reactivity pairs when elicited within a pure-list context.

**Experiment 1: Forward vs. Unrelated Pairs**

The goals of Experiment 1 were twofold. First, we sought to replicate positive reactivity on related pairs presented via mixed lists as initially reported by Soderstrom et al. (2015). We then tested whether this pattern would extend to pure lists by comparing participants who studied pure lists of forward associates to those who studied pure lists of unrelated pairs. Finally, across all list types, we included a group of participants who provided frequency judgments at encoding. Like JOLs, frequency judgments implicitly encourage participants to relate items together. However, this judgment type does not require participants to forecast subsequent memory and therefore is less likely to encourage metacognitive processes. Based on findings by Maxwell and Huff (in press), we expected that frequency judgment would produce reactivity patterns mirroring JOLs.

By comparing reactivity between mixed and pure lists, Experiment 1 provided a direct test of the changed-goal hypothesis while also providing a further test of the cue-strengthening account. First, because the changed-goal hypothesis states that reactivity results from participants changing study goals as they discern between related and unrelated pairs, reactivity would only be expected to occur when study pairs are presented using mixed lists, as pure lists lack the easy/difficult comparison required to trigger a goal change. The changed-goal hypothesis, therefore, predicts a null effect of reactivity for pairs presented using this list type, regardless of pair relatedness. The cue-strengthening account, however, makes no claims regarding easy/difficult comparisons. Instead, this account predicts positive reactivity for related pairs provided the encoding task emphasizes relatedness cues that are accessed at retrieval, regardless of whether participants study mixed or pure lists. Thus, if pure lists displayed the same reactivity patterns previously reported for mixed lists (i.e., positive reactivity for related pairs, no reactivity for unrelated pairs), this would provide further evidence for a cue-strengthening account rather than a goal changing account.

**Methods**

**Participants**

A total of 347 participants were recruited to complete Experiment 1. Participants were recruited from two sources: Undergraduate students from The University of Southern Mississippi’s undergraduate psychology research pool who completed the study in exchange for course credit (*n* = 260) and individuals who were recruited through Prolific Academic (www.prolific.co) who were compensated at a rate of $3.90/half hour (*n* = 87). Of these 347 participants, 111 were randomly assigned to the mixed-list group, which used a 3 × 2 mixed design in which pair relatedness was manipulated within subjects. The remaining 236 participants were randomly assigned to either the pure-related or unrelated-list groups, which employed a 3 × 2 between-subject design. For both groups, sample sizes were based on a set of a priori power analyses conducted with *G\*Power 3.1* (Faul, Erdfelder, Lang, & Buchner, 2007), which indicated that at least 42 participants would be needed to detect a medium effect with mixed lists (*d* = 0.50) and 158 participants would be needed to detect the same effect when analyzing pure lists. However, groups were oversampled due to an anticipated increase in participant performance variability via online data collection.

Within each list group, participants were further assigned to one of three groups based on encoding task (JOLs, frequency judgments, or silent reading/control). This resulted in a total of nine groups in (see Table 1 for each group’s final *n* following data screening). All participants were native English speakers who reported normal or corrected vision. Responses from 39 participants were excluded for one of the following reasons: (1) Low recall rates (e.g., correct recall rates < 5%) which suggested that participants did not correctly follow study instructions, or (2) recall rates of 100% across all blocks/pair types (which suggested participants were cheating during online testing). Additionally, data were omitted for one pure group participant due to a coding error. As a result, 307 participants were included in the following analyses (105 in the mixed-list analyses; 202 in the pure-list analyses). Final group *n*s are displayed in Table 1.

**Materials**

To create the stimuli, 200 word pairs were generated from the University of South Florida Free Association Norms (USF norms; Nelson, McEvoy, & Schreiber, 2004). These pairs were then divided into six study lists: Two mixed lists, two pure lists of forward pairs, and two pure lists of unrelated pairs. Mixed and pure list forward pairs were each matched on mean levels of forward associative strength (FAS) and backward associative strength (BAS). Additionally, all lists were matched on word length, SUBTLEX frequency values (Brysbaert & New, 2009), and concreteness values derived from the English Lexicon Project (Balota et al., 2007). Associative overlap measures and lexical characteristics for all stimuli are reported in Tables A1 and A2, respectively.

Study pairs across lists were randomized with the constraint that five non-tested buffer pairs were always presented at the beginning and end of each study list. All participants were presented with two study lists of the same type (i.e., participants in the pure unrelated condition would only receive the two pure unrelated study lists), which were organized into two study-test blocks. Block presentation order was counterbalanced across participants. Below, the process used to create the mixed and pure lists is described in further detail.

***Mixed Lists.*** To generate the mixed lists, 40 forward pairs (e.g., chisel-hammer) and 40 unrelated word pairs (e.g., justice-maroon) were randomly selected from the initial pool of 200 pairs. An additional 20 pairs (10 forward pairs and 10 unrelated pairs) were then selected as non-tested buffer items to control for primacy and recency effects. Pairs were divided into two study lists, each consisting of 20 forward and 20 unrelated study pairs and 10 buffer items (five related and five unrelated). As a result, each mixed list contained a total of 50 pairs.

***Pure Lists.*** Next, four pure lists were generated (two for each pair type). Starting with the related pure lists, each list contained 40 forward pairs, with list one consisting of the 40 pairs presented in the mixed list, and the other containing 40 forward pairs not assigned to a mixed list. The remaining 20 forward pairs served as primacy and recency buffers (10 per list). The second set of pure lists contained unrelated pairs and followed the same process used to create the related pure lists. Specifically, the first pure unrelated list used the 40 unrelated pairs presented in the mixed lists, while the second contained 40 unrelated pairs not assigned to a mixed list. Like the related lists, the remaining 20 unrelated pairs were used as buffer items. Thus, each pure list regardless of pair type contained of 40 study pairs and 10 buffer items.

**Procedure**

Data was collected online using *Collector*, an open-source program for presenting psychological experiments (Garcia & Kornell, 2015). Participants were randomly assigned to either the mixed- or pure-list groups and further randomly assigned to complete either the JOL, frequency judgment, or silent reading encoding tasks. Across groups, participants were informed that they would see a list of cue-target word pairs and that their memory for the target items in each pair would later be tested. Participants in the JOL and frequency judgment groups were further instructed to make judgments while encoding each study pair. Specifically, participants in the JOL group were instructed to rate the likelihood that they would be able to successfully recall the target item at test if prompted by only the cue. Participants in the frequency judgment group were instructed to rate the likelihood that the cue and target items would appear within the same context in natural language. Judgments in both groups utilized a 0-100 scale and were made concurrently with study, such that participants typed their ratings while the pair was displayed on the screen. Thus, the only difference between judgment conditions was the framing.

After receiving their respective encoding instructions, participants began the first study list. In the mixed-list groups, this list contained both forward and unrelated pairs. In contrast, participants assigned to the pure-list groups studied lists containing only forward or unrelated pairs. Following completion of the first study list, participants completed a 2 min filler task in which they listed the 50 U.S. states in alphabetical order. This was then followed by a cued-recall test which presented participants with each cue word from the preceding study list in a randomized order. Participants were instructed to type the correct target item from memory and were instructed to press the “Enter” key if they could not retrieve the correct item. Following completion of the cued-recall test, participants began the second block. This block followed the same format as the first, and participants studied the same list type in block 2 as block 1. Participants were debriefed following completion of the second block. The total experiment took approximately 30 minutes to complete.

**Results**

For all analyses, significance was set at *p* < .05. We report partial eta-squared (*ηp*2)and Cohen’s *d* effect sizes for all significant analyses of variance (ANOVAs) and *t*-tests. Additionally, all non-significant main effects, interactions, and post-hoc comparisons are supplemented by a Bayesian estimate of the strength of evidence in support of the null hypothesis (Masson, 2011; Wagenmakers, 2007). This analysis compares a model assuming a significant effect to a second model assuming a null effect. In doing so, a probability estimate can be generated, representing the likelihood that null hypothesis is retained (i.e., *p*BIC; Bayesian Information Criterion). Because this probably estimate is sensitive to sample size, it provides increased confidence in reported null effects.

The top panel of Figure 1 plots mean recall rates for participants who made JOLs, frequency judgments, or engaged in silent reading of mixed-list pairs, while the bottom panel displays mean recall rates between encoding groups for pure-list participants. For completeness, specific comparisons between forward and unrelated pairs are provided in Table A3.

**Mixed Lists**

First, a 2 (Pair Type: Forward vs. Unrelated) × 3 (Study Group: JOL vs. Frequency vs. No-JOL) mixed ANOVA was used to test for reactivity effects for pairs presented via mixed lists. A main effect of Pair Type was found, *F*(1, 102) = 1309.60, *MSE* = 99.84, *ηp*2 = .93, such that, when collapsed across encoding tasks, mean recall was higher for forward pairs (71.74) than unrelated pairs (21.69). The effect of Study Group was only marginally reliable, *F*(2, 102) = 2.64, *MSE* = 485.32, *p* = .08, *pBIC* = .88, but a significant interaction between Pair Type and Study Group was found, *F*(2, 102) = 12.41, *MSE* = 99.84, *ηp*2 = .20. Post-hoc *t*-tests indicated that for forward pairs, correct recall in both the JOL (75.59) and frequency judgment (76.68) groups exceeded that of the no-JOL group (62.98). All comparisons differed, *t*s ≥ 3.30, *d*s ≥ 0.78, except for the difference in recall between the JOL and frequency judgment groups, *t* < 1, *SEM* = 3.57, *p* *=* .74, *p*bic = .89. Importantly, for unrelated pairs, recall rates did not differ between the JOL (18.14) and frequency judgment (25.27) and the no-JOL (21.86) group, *t*s < 1, *p*s ≥ .38, *p*bics ≥ .85, though the comparison between the JOL and frequency judgment groups was marginal, *t*(68) = 1.91, *SEM* = 3.78, *p* *=* .06, *d* = 0.45, *p*bic = .58. Thus, when pairs were presented using mixed lists, JOL ratings and frequency judgments produced equivalent reactivity patterns for related pairs but no reactivity on unrelated pairs.

**Pure Lists**

A 2 (Pair Type: Forward vs Unrelated) × 3 (Study Group: JOL vs Frequency vs No-JOL) between-subject ANOVA tested whether reactivity patterns observed for mixed lists would hold for pairs in a pure-list context. Overall, this analysis yielded an effect of Pair Type, *F*(1, 196) = 468.13, *MSE* = 262.08, *ηp*2 = .70. Collapsed across encoding tasks, mean recall was higher for forward pairs (71.74) versus unrelated pairs (21.69). Next, a significant effect of Study Group emerged, *F*(2, 196) = 3.52, *MSE* = 262.08, *ηp*2 = .03, such that collapsed across pair type, mean recall was highest when participants made frequency judgments (50.69), followed by the JOL (51.40) and No-JOL groups (46.65). Post-hoc testing, however, revealed no significant differences in recall between encoding groups, *t*s < 1, *p*s ≥ .36, *p*bics ≥ .88.

Critically, a significant interaction emerged, *F*(2, 196) = 7.37, *MSE* = 262.08, *ηp*2 = .07. Follow-up testing revealed that for forward pairs, correct recall was greater in the JOL (83.19) and frequency judgment (77.78) groups relative to the no-JOL group (65.88). All comparisons differed significantly, *t*s ≥ 2.62, *d*s ≥ 0.70, except for the difference between the JOL and frequency judgment groups, *t*(60) = 1.36, *SEM* = 4.05, *p* *=* .18, *p*bic = .76. For unrelated pairs, correct recall did not differ between the between the JOL (23.25), frequency judgment (28.01), or the No-JOL (27.45) groups, *t*s ≤ 1.42, *p*s ≥ .16, *p*bic ≥ .76. Therefore, pure lists demonstrated similar reactivity patterns as mixed lists.

**Discussion**

The primary goal of Experiment 1 was to test the effect of list type on reactivity. In doing so, this experiment assessed reactivity effects for a group of participants who studied a mixed list of forward and unrelated pairs and tested whether these effects would extend to pairs presented in a pure-list context in which only one pair type was studied. Starting with participants in the mixed-list group, the predicted pattern of reactivity emerged. Relative to the control group, making JOLs increased correct recall of forward pairs—a positive reactivity pattern—but produced no recall benefit for unrelated pairs. This finding replicates previous work on JOL reactivity (e.g., Janes et al., 2018; Soderstrom et al. 2015). Finally, reactivity patterns observed for JOLs again extended to frequency judgments, replicating findings by Maxwell and Huff (in press) and further suggesting that JOL reactivity results from strengthening relational cues rather than via a metamemorial or predictive process.

Importantly, Experiment 1 showed that previously reported reactivity effects are not restricted to mixed-list designs. Pure lists showed positive JOL reactivity patterns for related pairs that mirrored mixed lists, and again, this reactivity pattern extended to frequency judgments. Because reactivity extended to pure lists, these effects are not simply the result of a comparison process (i.e., participants prioritizing easy pairs at the expense of more difficult ones as predicted by the changed-goal hypothesis). Instead, reactivity appears driven almost exclusively by pair relatedness, which further supports a cue-strengthening account (Soderstrom et al., 2015). The cue-strengthening account, however, also posits that for reactivity to occur, cues used to inform JOLs (e.g., relatedness) must be made available at test. For backward pairs (e.g., card-credit), the cue and target are related, yet the target item is an uncommon response to the cue. Thus, while backward pairs are thematically related, relatedness cues are not readily available at retrieval. As a result, it is unclear whether cue-strengthening can occur with backward pairs, given that the target item is a less obvious response to the cue.

To test this possibility, Experiment 2 compared mixed- and pure-list reactivity patterns using backward and unrelated pairs. Like forward pairs, participants assign backward pairs high JOL ratings at study (indicating that participants perceive backward pairs as related), but at test, participants often struggle to correctly retrieve the target (e.g., *the illusion of competence*; Koriat & Bjork, 2005). Backward pairs therefore provide a situation in which the cue-target word pair appears strongly related at encoding, but cues used to inform the judgment are not readily available at test. Finally, Experiment 2 similarly included a frequency judgment group, which tested whether JOL reactivity patterns would continue to extend to this encoding task in the absence of forward pairs.

**Experiment 2: Backward vs. Unrelated Pairs**

The goal of Experiment 2 was to test whether pure-list reactivity effects observed for forward pairs in Experiment 1 would extend to backward pairs. Like the previous experiment, Experiment 2 provided another test of the changed-goal and cue-strengthening accounts of reactivity. Based on the changed-goal hypothesis, positive reactivity would be expected to occur for backward pairs presented in a mixed list, given that this pair type appears easier to encode relative to unrelated pairs. However, no reactivity would be expected for pure lists, regardless of pair type. Regarding the cue-strengthening account, the presence of relatedness cues at encoding should boost recall of backward pairs compared to unrelated pairs, regardless of list type. However, because relatedness cues for backward pairs is not readily available at retrieval (i.e., the target is a less common response to the cue), any reactivity effects for backward pairs should be reduced compared forward pairs an Experiment 1. Finally, frequency judgments should again display reactivity patterns that mimic those found for JOLs in mixed and pure lists.

**Methods**

**Participants**

Experiment 2 followed the same design as Experiment 1. A separate set of 253 participants were recruited and completed the experiment online. Of these participants, 204 were undergraduate students from the University of Southern Mississippi who completed the study online in exchange for course credit. The remaining 49 participants were recruited via Prolific Academic and were paid $3.90 per half-hour of participation. Of the 253 participants recruited, 127 were randomly assigned to the mixed-list group, with the remaining 126 participants assigned to the pure related list group. Finally, the 106 participants who were assigned to the pure-unrelated group in Experiment 1 served as the pure unrelated comparison group. Thus, the pure-list groups contained a total of 232 participants. For both groups, sample sizes were based on Experiment 1. A sensitivity analysis conducted with *G\*Power 3.1* indicated that both the mixed and pure list samples were sufficient for detecting small-to-medium sized effects and interactions (*ds* = 0.26 and 0.40 for mixed and pure groups, respectively).

Like Experiment 1, participants in each list group were further assigned to randomly complete one of the three encoding tasks (JOLs, frequency judgments, or silent reading). Therefore, the following analyses include a total of nine groups (see Table 1 for final group *n*s following data screening). All participants were native English speakers reporting normal or corrected vision.

**Materials and Procedure**

Experiment 2 used the same study lists as the previous experiment, with the following modification. While the same unrelated word pairs from Experiment 1 were retained, all forward pairs (e.g., trout-fish) were replaced with backward pairs (e.g., fish-trout). Additionally, two pure lists containing only backward pairs were created, which provided a baseline for backward pair recall in the absence of unrelated study pairs. Study lists were identical to Experiment 1 in all other aspects including number of items, the inclusion of buffer pairs, and the study procedure (see Tables A1 and A4 for stimuli properties).

**Results**

Figure 2 (top panel) displays mean recall rates as a function of encoding group for participants who studied mixed lists. The bottom panel compares mean recall for each of the pure-list groups. For completeness, comparisons between pair types mixed and pure lists are provided in the Table A5. Data screening followed the same criteria used in Experiment 1, and across groups, responses from 13 participants were omitted. As a result, 120 participants were included in the mixed-list analyses, and 226 participants in the pure-list analyses (see Table 1 for final group ns).

**Mixed Lists**

A 2 (Pair Type: Backward vs. Unrelated) × 3 (Study Group: JOL vs. Frequency vs. No-JOL) mixed ANOVA was used to test for reactivity effects within mixed lists. This analysis yielded a main effect of Pair Type, *F*(1, 117) = 246.79, *MSE* = 87.63, *ηp*2 = .68. Collapsed across encoding groups, cued-recall was higher for backward pairs (43.90) than unrelated pairs (24.43). The main effect of Encoding Group, however, was non-significant *F*(2, 117) = 1.90, *MSE* = 600.55, *p* = .15, *pBIC* = .62, but the interaction was reliable, *F*(2, 117) = 15.83, *MSE* = 87.63, *ηp*2 = .22. Post-hoc testing confirmed the presence of positive reactivity for backward pairs, as recall was greatest for participants making frequency judgments (48.90), followed by participants in the JOL (46.84) and no-JOL groups (34.85). All comparisons differed significantly (*t*s ≥ 2.72, *d*s ≥ 0.62), except between the JOL and frequency judgment groups, *t* < 1, *p* = .66, *p*bic = .89. For unrelated pairs, recall rates were statistically equivalent between the frequency (26.75), JOL (20.98), and no-JOL groups (25.45; *t*s ≤ 1.68, *p*bics ≥ .69), indicating no reactivity. Reactivity patterns observed with forward pairs in mixed lists therefore extended to backward pairs.

**Pure Lists**

Next, a 2 (Pair Type: Backward vs. Unrelated) × 3 (Study Group: JOL vs. Frequency vs. No-JOL) between-subject ANOVA tested whether reactivity occurred for pairs presented within pure lists. Consistent with the previous analyses, a significant effect of pair type emerged, *F*(1, 220) = 42.91, *MSE* = 312.67, *ηp*2 = .16, such that recall of backward pairs (41.95) exceeded recall of unrelated pairs (26.25) when collapsing across encoding groups. However, the effect of Encoding Group was non-significant, *F*(2, 220) = 2.08, *MSE* = 312.67, *p* = .13, *p*bic = .65. Finally, the interaction between Pair Type and Encoding Group was right at the conventional level of significance, *F*(2, 220) = 2.95, *MSE* = 312.67, *p* = .05, *p*bic = .44, *ηp*2 = .03, and post-hoc comparisons were carried out as originally planned. Starting with backward pairs, correct recall was highest for participants in the frequency judgment group (46.01), followed by participants in the JOL (44.21), and no-JOL groups (34.83). Post-hoc *t*-tests confirmed that all comparisons differed significantly, *t*s ≥ 2.08, *d*s ≥ 0.47, except for the comparison between JOLs and frequency judgments, *t*(81) < 1, *SEM* = 4.39, *p* = .67, *p*bic = .89. Recall of unrelated pairs did not differ as a function of encoding group, *t*s ≤ 1.42, *p*s ≥ .16, *p*bic ≥ .76 (see Experiment 1). Thus, positive reactivity patterns observed for backward pairs in mixed lists extended to pure lists.

**Discussion**

Experiment 2 tested whether reactivity patterns observed for forward pairs in Experiment 1 would occur using backward pairs in which the target was less predictive of the cue at test. In doing so, this experiment provided an additional test of the cue-strengthening account of reactivity, as backward pairs provide a situation in which cues used to inform the JOL are less likely to be available at test. Furthermore, the inclusion of both mixed and pure lists allowed for an additional test of the changed-goal hypothesis. Overall, JOLs and frequency judgments each produced reactivity on backward pairs, regardless of list type. For unrelated pairs, however, no reactivity occurred. These findings are consistent with the previous experiments and provide additional support for the cue-strengthening account, as reactivity was again not limited to mixed in lists in which participants could distinguish between related and unrelated pairs.

In addition to providing additional tests of the changed-goal and cue-strengthening accounts of JOL reactivity, Experiment 2 provided a novel contribution to the reactivity literature by omitting the forward associate comparison group in favor of backward pairs. Studies investigating reactivity have largely focused on comparisons between forward and unrelated pairs (though see Maxwell & Huff, in press, and Mitchum et al., 2016 who each included backward pair comparison), and no reactivity study using related pairs has only targeted backward pairs without a forward pair comparison. Given the focus in the literature on forward associative pairs, Experiment 3 tested for reactivity on symmetrical associative pairs (e.g., king-queen) versus unrelated pairs. While backward pairs have been used in studies investigating the accuracy of JOLs (e.g., Koriat & Bjork, 2005), to date, little work on JOLs has used symmetrical pairs (cf. Maxwell & Huff, 2021). Furthermore, apart from Maxwell and Huff (in press), no study has investigated JOL reactivity effects using symmetrical paired associates. Experiment 3 therefore examined reactivity effects in mixed and pure lists using symmetrical pairs. In doing so, this experiment provided an additional opportunity to test whether reactivity effects would emerge on pure lists while further testing JOL reactivity accounts.

**Experiment 3: Symmetrical vs. Unrelated Pairs**

Experiment 3 tested whether JOL reactivity would extend to symmetrical pairs (e.g., salt-pepper) when presented in mixed lists with unrelated pairs or when presented in isolation via pure lists. Like backward pairs, symmetrical pairs can be deceptive as they contain strong backward associations. However, these pairs also contain strong forward associations, which should make them easier to learn relative to backward pairs (Maxwell & Huff, 2021). The use of symmetrical pairs in Experiment 3 is important, as it provides a novel pair type with which to test for reactivity effects. Therefore, our use of symmetrical pairs provides a further test of the changed-goal and cue-strengthening accounts while also testing the generality of JOL reactivity effects. Based on the previous experiments, findings were expected to conform to a cue-strengthening account, with positive reactivity occurring for symmetrical pairs and no reactivity for unrelated pairs. Furthermore, this pattern was expected to occur regardless of whether participants studied mixed or pure lists. Finally, frequency judgments were again expected to produce reactivity patterns mirroring JOLs.

**Methods**

**Participants**

A total of 227 participants were recruited to complete Experiment 3. Like the previous experiments, participants were either undergraduates recruited from the University of Southern Mississippi’s psychology research pool (*n* = 187) who completed the study online in exchange for course credit or individuals recruited through Prolific Academic who completed the study online at a rate of $3.90/half hour (*n* = 40). Of these participants, 113 were randomly assigned to the mixed-list group, with the remainder randomly assigned to the pure-symmetrical group (*n* = 114). Like Experiment 2, the 106 participants who studied pure unrelated lists in Experiment 1 again served as the pure unrelated comparison group. Therefore, the pure-list group contained a total of 220 participants. Group sizes were informed by the sample used in Experiment 1, and a sensitivity analysis via *G\*Power 3.1* confirmed that the mixed- and pure-list groups were sufficient for detecting small-to-medium main effects and interactions (*ds* ≥ 0.42). Like the preceding experiments, participants within both list groups were further randomly assigned to either the JOL, frequency, or no-JOL encoding groups. Nine groups are included in the following analyses (see Table 1 for final group *n*s after data screening).

**Materials and Procedure**

Experiment 3 used a modified version of the study lists presented in Experiments 1 and 2. While the same unrelated word pairs from the previous experiments were retained, the forward/backward pairs were replaced with symmetrical pairs (e.g., king-queen). Unlike forward and backward pairs which are characterized by an asymmetrical associative relationship (i.e., from cue to target in forward pairs or vice-versa in backward pairs), symmetrical pairs contain relationships in both directions of similar associative strength. All other aspects of the study lists as well as the study procedure were identical to Experiments 1 and 2 (see Tables A1 and A6 for stimuli properties).

**Results**

Figure 3 (top panel) shows recall rates for participants who studied mixed lists as a function of encoding task, while the bottom panel displays mean recall rates for each encoding task across pure list groups. For completeness, all comparisons between related and unrelated pairs are provided in the Appendix (Table A7). Data screening followed the same procedure outlined in Experiment 2, and data from 18 participants were omitted (see Table 1 for group *n*s).

**Mixed Lists**

Like the previous experiments, a 2 (Pair Type: Symmetrical vs. Unrelated) × 3 (Study Group: JOL vs. Frequency vs. No-JOL) mixed ANOVA was used to test for reactivity effects in mixed lists. An effect of Pair Type was found, *F*(1, 103) = 825.46, *MSE* = 112.87, *ηp*2 = .89, as recall of symmetrical pairs (65.09) exceeded recall of unrelated pairs (23.17). The effect of Encoding Group was non-significant, *F*(2, 103) = 1.33, *MSE* = 497.13, *p* = .27, *p*bic = .96, but an interaction was found, confirming the presence of a reactivity pattern, *F*(2, 103) = 12.57, *MSE* = 112.87, *ηp*2 = .20. For symmetrical pairs, mean recall was highest following frequency judgments (69.34), than JOLs (69.33) and the no-JOL control task (56.51). Follow up *t*-tests confirmed that all comparisons differed significantly (*t*s ≥ 2.78, *d*s ≥ 0.65), except for the comparison between frequency judgments and JOLs, *t* < 1, *SEM* = 3.88, *p* = .99, *p*bic = .99. For unrelated pairs, no reactivity was observed. Mean recall did not differ between the JOL (21.24), frequency (23.46), or no-JOL encoding groups (24.80; *t*s < 1, *p*s ≥ .40, *p*bics ≥ .85). Thus, mixed list reactivity patterns with forward and backward paired associates extend to symmetrical pairs.

**Pure Lists**

A 2 (Pair Type: Symmetrical vs. Unrelated) × 3 (Study Group: JOL vs. Frequency vs. No-JOL) between-subject ANOVA was then used to test reactivity effects for symmetrical pairs would extend to pure lists. Consistent with the previous experiments, this analysis yielded a significant effect of Pair Type, *F*(1, 203) = 407.21, *MSE* = 246.60, *ηp*2 = .67. Across encoding groups, recall of symmetrical pairs (70.08) was greater than unrelated pairs (26.25). Additionally, significant effect of Encoding Group was detected, *F*(2, 203) = 6.84, *MSE* = 246.60, *ηp*2 = .06, such that recall was highest for participants in the frequency judgment group (52.57), followed by the JOL (47.31) and no-JOL groups (43.39). Post-hoc tests, however, indicated that this effect was driven by difference between the frequency judgment and no-JOL groups, *t*(140) = 2.09, *SEM* = 4.44, *p* = .04, *d* = 0.35. All other comparisons were non-significant, *t*s ≤ 1.06, *p*s ≥ .29, *p*bics ≥ .90. Importantly, a significant interaction was again found, *F*(2, 203) = 8.12, *MSE* = 246.60, *ηp*2 = .07. For symmetrical pairs, recall was highest for participants in the frequency judgment group (77.81), followed by the JOL (73.63) and no-JOL groups (58.89). All comparisons differed significantly, *t*s ≥ 3.80, *d*s ≥ 0.85, apart from the comparison between the JOL and frequency groups, *t*(66) = 1.12, *SEM* = 3.81, *p* = .26, *p*bic = .81. For unrelated pairs, recall again did not differ between encoding groups, *t*s ≤ 1.42, *p*s ≥ .16, *p*bic ≥ .76 (see Experiment 1). Thus, like the previous experiments, JOLs and frequency judgments again produced a positive reactivity effect on related pairs in a pure list setting.

**Discussion**

The goal of Experiment 3 was to test whether reactivity effects observed for forward and backward pairs in Experiments 1 and 2 would extend to symmetrical pairs. Overall, both JOLs and frequency judgments produced positive reactivity effects on symmetrical pairs, and as observed in the previous experiments, neither judgment type produced a reactive effect on unrelated pairs. Importantly, reactivity on symmetrical pairs occurred regardless of whether participants studied mixed or pure lists, further suggesting that reactivity is not due to context in which items are studied (i.e., easy/related vs. difficult/unrelated study materials) as posited by the changed-goal hypothesis. Therefore, findings from Experiment 3 align with our previous experiments while providing additional support for a cue-strengthening account.

**General Discussion**

The present study tested the changed-goal and cue strengthening accounts of JOL reactivity by investigating whether reactivity patterns previously reported on mixed lists (i.e., positive reactivity on related pairs, no reactivity on unrelated pairs; Janes et al., 2018; Maxwell & Huff, in press; Soderstrom et al., 2015) would emerge when pairs were presented in isolation via pure lists. In doing so, each experiment focused exclusively on one type of related paired associate type (forward, backward, or symmetrical) and directly compared it to unrelated pairs within both mixed- and pure-list contexts. A secondary goal was to further test whether reactivity effects were unique to JOLs. Therefore, in addition to the JOL versus no-JOL comparison traditionally used to explore reactivity, each experiment also included a group of participants who completed a frequency judgment task in lieu of providing JOLs. This additional comparison group was included to evaluate whether any observed reactivity patterns would occur when a non-metacognitive judgment task was used.

Overall, Experiment 1 replicated previous JOL reactivity patterns for mixed lists (e.g., Janes et al., 2018, Maxwell & Huff, in press, Soderstrom et al., 2015), such that JOLs produced positive reactivity on forward pairs but were not reactive on unrelated pairs. Importantly, this reactivity pattern extended to pure lists in all three experiments, suggesting that reactivity is not driven by changes in participant study goals. Additionally, all observed reactivity on JOLs extended to frequency judgments, providing additional evidence that reactivity effects are driven by the encoding task strengthening relatedness cues used at retrieval rather than via a comparative process as posited by the changed-goal hypothesis. Experiments 2 and 3 then showed that these positive reactivity patterns for both JOLs and frequency judgments extend to backward and symmetrical paired associates, respectively. Thus, across experiments and list types, negative reactivity for unrelated pairs as reported by Mitchum et al. (2016) consistently failed to occur. Therefore, a key finding from the present study is that JOLs consistently produce positive reactivity on related pairs but no reactivity on unrelated pairs, regardless of the context in which pairs are presented.

The finding that positive reactivity extends to related pairs in pure lists provides important insights regarding JOL reactivity effects. Regarding the changed-goal hypothesis, Mitchum et al. (2016) proposed that reactivity occurs as a byproduct of participants altering their study goals as a function of pair difficulty (i.e., metacognitive evaluation of pair difficulty produces a shift in study goals). However, this account cannot explain reactivity effects in pure lists, given that pure lists lack the easy/difficult comparison necessary to trigger a shift in study goals. Therefore, our pure-list reactivity findings are inconsistent with a changed-goal account. Concerning Soderstrom et al.’s (2015) cue-strengthening account, the extension of reactivity patterns to pure lists further supports the notion that reactivity is driven by relational encoding that is selectively applied to related but not unrelated pairs. As such, pure list reactivity findings observed in the present study are in-line with this account.

In addition to testing for reactivity effects between list types, each experiment included an additional comparison group in which participants rated the likelihood of words co-occurring together. We included these groups to test whether reactivity patterns observed on non-metacognitive judgment in mixed lists reported by Maxwell and Huff (in press) would similarly extend to pure lists. Like JOLs, frequency judgments direct attention towards relational aspects of study pairs without explicitly instructing participants to employ a relational strategy at encoding. Additionally, this task used the same 0-100 rating scale as JOLs. Thus, the frequency judgment task resembled JOLs, but removed the requirement that participants forecast later recall. Across experiments, frequency judgments consistently showed reactivity patterns mirroring JOLs, such that frequency that these judgments provided a memory boost to related pairs but no reactivity when pairs were unrelated. Furthermore, like JOLs, frequency judgments were reactive regardless of whether participants studied pairs within mixed or pure lists. Thus, metacognitive processes induced by JOLs do not appear to be a requisite for reactivity to occur.

While our comparison of mixed versus pure lists was designed to test the changed-goal and cue strengthening accounts of reactivity, we note that the present study may also provide insight regarding participant strategy use. First, our finding that JOL reactivity extends to frequency judgments replicates previous work by Maxwell & Huff (in press). To explain this observation, Maxwell & Huff proposed that JOLs implicitly encourage participants to relate study pairs together at encoding. However, this relational encoding is applied strategically, such that only related pairs receive a memory benefit. Within this context, our finding that both JOL and frequency judgments are reactive on related pairs presented in pure lists may qualify this strategy use account, indicating that participants may be able to apply a relational encoding strategy in both mixed and pure list contexts. The pure list pattern is important because it indicates that participants do not need to be exposed to unrelated pairs to instantiate relational encoding. Finally, we note that in addition to testing the cue-strengthening account, Rivers et al. (2021) also assessed strategy use by having participants report the encoding strategies used on each pair following retrieval of each target. Reported strategies did not differ between related and unrelated pairs; however, because strategy use was assessed at retrieval, this measure did not capture online strategy use at encoding. More work is therefore needed to fully understand the role that strategy use plays in JOL reactivity including whether strategies are shifted across items within a study set.

While the present study replicated previous work showing positive reactivity on related pairs, we note that for each experiment, participant study was self-paced. Although other studies investigating reactivity have also made use of self-paced study (e.g., Janes et al., 2018; Mitchum et al., 2016; see Maxwell & Huff, in press, for a review), the memory improvements observed for both JOLs and frequency judgments could potentially be attributed to participants in the judgment groups encoding pairs for longer durations relative to the silent reading group. However, across experiments and list types, encoding durations were generally longer for participants in the control groups compared to the judgment groups (see Tables A8-A9). Thus, the reactivity effects observed in the present study do not appear to be driven by longer encoding durations and instead likely reflect additional processing due to making judgments at encoding.

**Conclusion**

Researchers have become increasingly interested in the reactive effects of immediate JOLs on cue-target word pairs. The present study tested the changed-goal and cue-strengthening accounts of reactivity by testing between mixed (e.g., related and unrelated pairs) and pure study lists (e.g., only unrelated pairs). Across three experiments, we show that JOLs produced positive reactivity on related pairs (forward, backward, or symmetrical paired associates) but were not reactive on unrelated pairs, regardless of whether participants studied pairs within mixed or pure-list contexts. Additionally, we replicate previous findings showing that JOL reactivity extends to other, non-metacognitive judgment tasks. As a result, the present study provides further evidence for a cue-strengthening account of JOL reactivity rather than a goal-changing account.

**Open Practices Statement**

The data for all experiments have been made available at https://osf.io/3fztn/. None of the experiments were preregistered.

References

Akdoğan, E., Izaute, M., Danion, J., Vidailhet, P., & Bacon, E. (2016). Is retrieval the key? Metamemory judgment and testing as learning strategies. *Memory, 24*(10), 1390-1395.

Arbuckle, T. Y., & Cuddy, L. L. (1969). Discrimination of item strength at time of presentation. *Journal of Experimental Psychology*, *81* (1), 126–131.

Balota, D. A., Yap, M. J., Hutchison, K. A., Cortese, M. J., Kessler, B., Loftis, B., Neely, J. H., Nelson, D. L., Simpson, G. B, & Treiman, R. (2007). The English lexicon project. *Behavior Research Methods, 39* (3), 445-459.

Brysbaert, M., & New, B. (2009). Moving beyond Kučera and Francis: A critical evaluation of current word frequency norms and the introduction of a new and improved word frequency measure for American English. *Behavior Research Methods, 41*, 977–990.

Double, K. S., Birney, D. P., & Walker, S. A. (2018). A meta-analysis and systematic review of reactivity to judgments of learning. *Memory, 26* (6), 741-750.

Ericsson, K. A., & Simon, H. A. (1993). *Protocol analysis: Verbal reports as data (Rev. ed.)*. Cambridge, MA: Bradford Books/ MIT Press.

Faul, F., Erdfelder, E., Lang, A. G., & Buchner, A. (2007). G\*Power 3: a flexible statistical power analysis program for the social, behavioral, and biomedical sciences. *Behavior* *Research Methods*, *39* (2), 175–191.

Garcia, M. & Kornell, N. (2015). Collector [Computer software]. Retrieved April 3rd, 2020 from https://github.com/gikeymarica/Collector.

Janes, J. L., Rivers, M. L, & Dunlosky, J. (2018). The influence of making judgments of learning on memory performance: Positive, negative, or both? *Psychonomic Bulletin & Review, 25* (6), 2356-2364.

Koriat, A. (1997). Monitoring one’s own knowledge during study: A cue-utilization approach to judgments of learning. *Journal of Experiment Psychology: General, 126* (4), 349-370.

Koriat, A., & Bjork, R. A. (2005). Illusions of competence in monitoring one’s knowledge during study. *Journal of Experimental Psychology: Learning, Memory, and Cognition*, *31* (2), 187–194.

Koriat, A., Sheffer, L., & May’ayan, H. (2002). Comparing objective and subjective learning curvs: Judgments of learning exhibit increased underconfidence with practice. *Journal of Experimental Psychology: General, 131*(2), 147-162.

Luna, K., Albuquerque, P. B., & Martín-Luengo, B. (2019). Cognitive load eliminates the effect of perceptual information on judgments of learning with sentences. *Memory & Cognition, 47*, 106-116.

Maki, W. S. (2007). Judgments of associative memory. *Cognitive Psychology, 54*(4), 319-353.

Masson, M. E. J. (2011). A tutorial on a practical Bayesian alternative to null-hypothesis significance testing. *Behavior Research Methods, 43*, 679-690.

Maxwell, N. P., & Huff, M. J. (2021). The deceptive nature of associative word pairs: Effects of associative direction on judgments of learning. *Psychological Research, 85*, 1757-1775.

Maxwell, N. P., & Huff, M. J. (in press). Reactivity from judgments of learning is not only due to memory forecasting: Evidence from associative memory and frequency judgments. *Metacognition and Learning*, 1-37.

Meeter, M., & Nelson, T. O. (2003). Multiple study trials and judgments of learning. *Acta Psychologica, 113*(2), 123-132.

Myers, S. J., Rhodes, M. G., & Hausman, H. E. (2020). Judgments of learning (JOLs) selectively improve memory depending on the type of test. *Memory & Cognition, 48*, 745-758.

Mitchum, A. L., Kelley, C. M., & Fox, M. C. (2016). When asking the question changes the ultimate answer: Metamemory judgments change memory. *Journal of Experimental Psychology: General, 145* (2), 200-219.

Nelson, D. L., McEvoy, C. L., & Schreiber, T. A. (2004). The University of South Florida free association, rhyme, and word fragment norms. *Behavior Research Methods,* *Instruments, & Computers*, *36* (3), 402–407.

Nelson, T. O. & Narens, L. (1990). Metamemory: A theoretical framework and new findings. In: *The psychology of learning and motivation*, ed. G. Bower. American Psychologist.

Rhodes, M. G. (2016). Judgments of learning. In J. Dunlosky and S. K. Tauber (Eds.), *The Oxford Handbook of Metamemory* (pp. 65-80). New York: Oxford University Press.

Rhodes, M. G, & Tauber, S. K. (2011). The influence of delaying judgments of learning on metacognitive accuracy: A meta-analytics review. *Psychological Bulletin, 137*(1), 131-148.

Rivers, M. L., Janes, J. L., & Dunlosky, J. (2021). Investigating memory reactivity with a within-participant manipulation of judgments of learning: Support for the cue-strengthening hypothesis. *Memory, 29*(10), 1342-1353.

Soderstrom, N. C., Clark, C. T., Halamish, V., & Bjork, E. L. (2015). Judgments of learning as memory modifiers. *Journal of Experimental Psychology: Learning, Memory, and Cognition, 41*, 553–558.

Spellman, B. A., & Bjork, R. A. (1992). When predictions create reality: Judgments of learning may alter what they are intended to assess. *Psychological Science, 3*(5), 315-317.

Tauber, S. K., & Witherby, A. E. (2019). Do judgments of learning modify older adults’ actual learning? *Psychology and Aging, 34*(6), 836-847.

Townsend, C. L., & Heit, E. (2011). Judgments of learning and improvement. *Memory & Cognition, 39*, 204-216.

Valentine, K. D., & Buchanan, E. M. (2013). JAM-boree: An application of observation oriented modeling to judgements of associative memory. *Journal of Cognitive Psychology, 25*(4), 400-422.

Wagenmakers, E. (2007). A practical solution to the pervasive problems of *p* values. *Psychonomic Bulletin & Review, 14*, 779-804.

Table 1

*Final Sample Sizes for all Comparison Groups in each Experiment.*

|  |  |  |  |  |  |  |  |
| --- | --- | --- | --- | --- | --- | --- | --- |
| Experiment | Encoding Task | Mixed | Pure Forward | Pure Backward | Pure Symmetrical | Pure Unrelated | |
| Exp. 1 | JOL | 36 | 31 | -- | -- | 35 |
|  | Frequency | 34 | 31 | -- | -- | 37 |
|  | No-JOL | 35 | 34 | -- | -- | 34 |
| Exp. 2 | JOL | 40 | -- | 41 | -- | 35 |
|  | Frequency | 43 | -- | 42 | -- | 37 |
|  | No-JOL | 37 | -- | 37 | -- | 34 |
| Exp. 3 | JOL | 35 | -- | -- | 32 | 35 |
|  | Frequency | 36 | -- | -- | 36 | 37 |
|  | No-JOL | 35 | -- | -- | 35 | 34 |

*Note*: Cells reflect final *n*s for each group following data screening. The five left-most columns denote list type. The pure unrelated group in Experiment 1 was used as the pure unrelated comparison in Experiments 2 and 3.



*Figure 1.* Mean percent recall for participants in Experiment 1 who completed the JOL, frequency judgment, or No-JOL silent reading tasks for mixed lists (top panel) or pure lists (bottom panel). Error bars represent 95% confidence intervals.



*Figure 2.* Mean percent recall for participants in Experiment 2 who completed the JOL, frequency judgment, or No-JOL silent reading tasks for mixed lists (top panel) or pure lists (bottom panel). Error bars represent 95% confidence intervals.



*Figure 3.* Mean percent recall for participants in Experiment 3 who completed the JOL, frequency judgment, or No-JOL silent reading tasks for mixed lists (top panel) or pure lists (bottom panel). Error bars represent 95% confidence intervals.

**Appendix**

Table A1

|  |  |  |  |  |  |  |
| --- | --- | --- | --- | --- | --- | --- |
|  | Pair Type | Variable | *M* | *SD* | *Min.* | *Max.* |
| Experiment 1 | Pure Forward | FAS | .37 | .21 | .05 | .81 |
|  |  | BAS | 0 | 0 | 0 | 0 |
|  | Mixed Forward | FAS | .37 | .21 | .05 | .81 |
|  |  | BAS | 0 | 0 | 0 | 0 |
| Experiment 2 | Pure Backward | FAS | 0 | 0 | 0 | 0 |
|  |  | BAS | .37 | .21 | .05 | .81 |
|  | Mixed Backward | FAS | 0 | 0 | 0 | 0 |
|  |  | BAS | .37 | .21 | .05 | .81 |
| Experiment 3 | Pure Symmetrical | FAS | .27 | .18 | .01 | .59 |
|  |  | BAS | .27 | .17 | .01 | .58 |
|  | Mixed Symmetrical | FAS | .19 | .13 | .01 | .46 |
|  |  | BAS | .19 | .13 | .02 | .52 |

*Summary Statistics for Associative Overlap Variables across each Experiment.*

*Notes.* Values are grouped by JOL condition. FAS and BAS values for unrelated pairs are not included as by deﬁnition these associations between these items have not been normed. Mean FAS and BAS values are computed by taking the average association strength for each pair.

Table A2

*Summary Statistics for Cue and Target Item Properties in Experiment 1*

|  |  |  |  |  |
| --- | --- | --- | --- | --- |
| Pair Type | Position | Variable | *M* | *SD* |
| Mixed Forward | Cue | Concreteness | 5.04 | 1.15 |
|  |  | Length | 5.83 | 1.89 |
|  |  | Frequency | 2.57 | 0.77 |
|  | Target | Concreteness | 4.94 | 1.11 |
|  |  | Length | 4.48 | 1.24 |
|  |  | Frequency | 3.72 | 0.65 |
| Mixed Unrelated | Cue | Concreteness | 3.94 | 3.91 |
|  |  | Length | 5.20 | 1.67 |
|  |  | Frequency | 3.79 | 1.41 |
|  | Target | Concreteness | 3.92 | 1.56 |
|  |  | Length | 5.22 | 1.37 |
|  |  | Frequency | 3.83 | 1.30 |
| Pure Forward | Cue | Concreteness | 4.81 | 1.00 |
|  |  | Length | 5.85 | 1.63 |
|  |  | Frequency | 2.49 | 0.65 |
|  | Target | Concreteness | 4.88 | 1.07 |
|  |  | Length | 4.48 | 1.38 |
|  |  | Frequency | 3.73 | 0.63 |
| Pure Unrelated | Cue | Concreteness | 4.52 | 1.26 |
|  |  | Length | 5.11 | 1.48 |
|  |  | Frequency | 3.05 | 0.84 |
|  | Target | Concreteness | 4.64 | 1.29 |
|  |  | Length | 5.08 | 1.34 |
|  |  | Frequency | 3.05 | 0.81 |

*Notes.* Values are grouped by list condition. Frequency is measured using SUBTLEX word frequency measure (Brysbaert & New, 2009). Concreteness and length were taken from the English Lexicon Project (Balota et al., 2007).

Table A3

*Comparisons of Mean Recall Percentages for each Encoding Task as a function of Pair Type in Experiment 1.*

|  |  |  |  |  |  |  |
| --- | --- | --- | --- | --- | --- | --- |
| Encoding Task | | List Type | Pair Type | *M* | *± 95% CI* | U |
| Mixed | JOL | | Forward | 75.59 | 4.63 | 4.34\* |
|  |  | | Unrelated | 18.14 | 3.99 |  |
|  | Frequency | | Forward | 76.68 | 5.11 | 3.05\* |
|  |  | | Unrelated | 25.27 | 6.18 |  |
|  | No-JOL | | Forward | 62.98 | 6.01 | 2.00\* |
|  |  | | Unrelated | 21.86 | 7.50 |  |
| Pure | JOL | | Forward | 83.19 | 2.56 | 4.66\* |
|  |  | | Unrelated | 23.25 | 3.56 |  |
|  | Frequency | | Forward | 77.78 | 4.60 | 2.96\* |
|  |  | | Unrelated | 28.01 | 3.27 |  |
|  | No-JOL | | Forward | 65.88 | 4.11 | 2.08\* |
|  |  | | Unrelated | 27.43 | 4.66 |  |

*Note.* The right-most column indicates Cohen’s *d* effect sizes for Related-Unrelated comparisons, \* = *p* < .05. U = Unrelated pairs.

Table A4

*Summary Statistics for Cue and Target Item Properties in Experiment 2*

|  |  |  |  |  |
| --- | --- | --- | --- | --- |
| Pair Type | Position | Variable | *M* | *SD* |
| Mixed Backward | Cue | Concreteness | 5.13 | 1.06 |
|  |  | Length | 4.48 | 1.24 |
|  |  | Frequency | 3.72 | 0.65 |
|  | Target | Concreteness | 4.82 | 1.17 |
|  |  | Length | 5.83 | 1.89 |
|  |  | Frequency | 2.57 | 0.77 |
| Mixed Unrelated | Cue | Concreteness | 4.73 | 1.23 |
|  |  | Length | 5.20 | 1.67 |
|  |  | Frequency | 3.19 | 0.93 |
|  | Target | Concreteness | 4.54 | 1.33 |
|  |  | Length | 5.23 | 1.37 |
|  |  | Frequency | 3.18 | 0.76 |
| Pure Backward | Cue | Concreteness | 5.03 | 1.13 |
|  |  | Length | 4.45 | 1.27 |
|  |  | Frequency | 3.75 | 0.62 |
|  | Target | Concreteness | 4.88 | 1.22 |
|  |  | Length | 6.17 | 1.86 |
|  |  | Frequency | 2.48 | 0.67 |

*Notes.* Values are grouped by list condition. Frequency is measured using SUBTLEX word frequency measure (Brysbaert & New, 2009). Concreteness and length were taken from the English Lexicon Project (Balota et al., 2007).

Table A5

*Comparisons of Mean Recall Percentages for each Encoding Task as a function of List and Pair Type in Experiment 2.*

|  |  |  |  |  |  |  |
| --- | --- | --- | --- | --- | --- | --- |
| Encoding Task | | List Type | Pair Type | *M* | *± 95% CI* | U |
| Mixed | JOL | | Backward | 46.84 | 6.07 | 1.47\* |
|  |  | | Unrelated | 20.99 | 4.79 |  |
|  | Frequency | | Backward | 48.90 | 6.20 | 1.18\* |
|  |  | | Unrelated | 26.75 | 4.97 |  |
|  | No-JOL | | Backward | 34.85 | 5.96 | 0.49\* |
|  |  | | Unrelated | 25.45 | 6.47 |  |
| Pure | JOL | | Backward | 44.21 | 4.96 | 1.17\* |
|  |  | | Unrelated | 23.25 | 3.32 |  |
|  | Frequency | | Backward | 46.01 | 3.76 | 1.16\* |
|  |  | | Unrelated | 28.01 | 3.04 |  |
|  | No-JOL | | Backward | 34.83 | 3.97 | 0.40 |
|  |  | | Unrelated | 27.43 | 4.46 |  |

*Note.* The right-most column indicates Cohen’s *d* effect sizes for Related-Unrelated comparisons, \* = *p* < .05. U = Unrelated pairs. Pure unrelated comparison is taken from Experiment 1.

Table A6

*Summary Statistics for Cue and Target Item Properties in Experiment 3*

|  |  |  |  |  |
| --- | --- | --- | --- | --- |
| Pair Type | Position | Variable | *M* | *SD* |
| Mixed Symmetrical | Cue | Concreteness | 4.70 | 1.38 |
|  |  | Length | 5.21 | 1.94 |
|  |  | Frequency | 3.23 | 0.67 |
|  | Target | Concreteness | 4.70 | 1.38 |
|  |  | Length | 5.21 | 1.94 |
|  |  | Frequency | 3.23 | 0.67 |
| Mixed Unrelated | Cue | Concreteness | 4.73 | 1.23 |
|  |  | Length | 5.20 | 1.67 |
|  |  | Frequency | 3.19 | 0.93 |
|  | Target | Concreteness | 4.54 | 1.33 |
|  |  | Length | 5.23 | 1.37 |
|  |  | Frequency | 3.18 | 0.76 |
| Pure Symmetrical | Cue | Concreteness | 4.63 | 1.41 |
|  |  | Length | 5.31 | 1.67 |
|  |  | Frequency | 3.24 | 0.74 |
|  | Target | Concreteness | 4.68 | 1.39 |
|  |  | Length | 5.16 | 1.76 |
|  |  | Frequency | 3.17 | 0.71 |

*Notes.* Values are grouped by list condition. Frequency is measured using SUBTLEX word frequency measure (Brysbaert & New, 2009). Concreteness and length were taken from the English Lexicon Project (Balota et al., 2007).

Table A7

*Comparisons of Mean Recall Percentages for each Encoding Task as a function of List and Pair Type in Experiment 3.*

|  |  |  |  |  |  |  |
| --- | --- | --- | --- | --- | --- | --- |
| Encoding Task | | List Type | Pair Type | *M* | *± 95% CI* | U |
| Mixed | JOL | | Symmetrical | 69.33 | 4.60 | 3.21\* |
|  |  | | Unrelated | 21.24 | 5.30 |  |
|  | Frequency | | Symmetrical | 69.34 | 5.86 | 2.76\* |
|  |  | | Unrelated | 23.46 | 4.97 |  |
|  | No-JOL | | Symmetrical | 56.51 | 7.02 | 1.56\* |
|  |  | | Unrelated | 24.80 | 6.47 |  |
| Pure | JOL | | Symmetrical | 73.63 | 4.04 | 3.18\* |
|  |  | | Unrelated | 23.25 | 3.53 |  |
|  | Frequency | | Symmetrical | 77.81 | 3.20 | 3.59\* |
|  |  | | Unrelated | 28.01 | 3.16 |  |
|  | No-JOL | | Symmetrical | 58.89 | 3.51 | 1.81\* |
|  |  | | Unrelated | 27.42 | 4.62 |  |

*Note.* The right-most column indicates Cohen’s *d* effect sizes for Related-Unrelated comparisons, \* = *p* < .05. U = Unrelated pairs. Pure unrelated comparison is taken from Experiment 1.

Table A8

Mean Encoding Latencies as a Function of Pair Type and Encoding Task for Mixed Lists in Experiments 1-3.

|  |  |  |  |  |  |
| --- | --- | --- | --- | --- | --- |
| Experiment | Encoding Task | Forward | Backward | Symmetrical | Unrelated |
| Exp. 1 | JOL | 4166 | -- | -- | 5009 |
|  | Frequency | 4500 | -- | -- | 5992 |
|  | Read | 6268 | -- | -- | 8150 |
| Exp. 2 | JOL | -- | 5527 | -- | 4995 |
|  | Frequency | -- | 5444 | -- | 5179 |
|  | Read | -- | 5396 | -- | 5801 |
| Exp. 3 | JOL | -- | -- | 5316 | 6470 |
|  | Frequency | -- | -- | 4322 | 5310 |
|  | Read | -- | -- | 5603 | 7103 |

Note: Cells display mean RTs in ms.

Table A9

*Mean Encoding Latencies as Functions of Pair Type and Encoding Tasks for Pure Lists in Experiments 1-3.*

|  |  |  |  |  |  |
| --- | --- | --- | --- | --- | --- |
| Experiment | Encoding Task | Forward | Backward | Symmetrical | Unrelated |
| Exp. 1 | JOL | 3483 | -- | -- | 5197 |
|  | Frequency | 3616 | -- | -- | 6407 |
|  | Read | 5249 | -- | -- | 6376 |
| Exp. 2 | JOL | -- | 6398 | -- | 5197 |
|  | Frequency | -- | 5743 | -- | 6407 |
|  | Read | -- | 6561 | -- | 6376 |
| Exp. 3 | JOL | -- | -- | 5026 | 5197 |
|  | Frequency | -- | -- | 4294 | 6407 |
|  | Read | -- | -- | 4739 | 6376 |

*Note:* Cells display mean RTs in ms. Pure unrelated comparison is taken from Experiment 1.